

Package ‘LAM’

May 11, 2017

Type Package

Title Some Latent Variable Models

Version 0.0-17

Date 2017-05-11 14:07:52

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Description Contains some procedures for latent variable modelling with a particular focus on multilevel data.

The LAM package contains mean and covariance structure modelling for multivariate normally distributed data ('mlnormal'), a general Metropolis-Hastings algorithm ('amh') and penalized maximum likelihood estimation ('pmle').

Depends R (>= 3.1)

Imports CDM, coda, graphics, Rcpp, sirt (>= 2.0), stats, utils

Suggests lavaan, lme4, MASS

LinkingTo Rcpp, RcppArmadillo

License GPL (>= 2)

NeedsCompilation yes

Repository CRAN

Date/Publication 2017-05-11 16:30:10 UTC

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Description

Contains some procedures for latent variable modelling with a particular focus on multilevel data. The LAM package contains mean and covariance structure modelling for multivariate normally distributed data ('mlnormal'), a general Metropolis-Hastings algorithm ('amh') and penalized maximum likelihood estimation ('pmle').

Details

The **LAM** package contains the following main functions:

- A general fitting method for mean and covariance structure for multivariate normally distributed data is the `mlnormal` function. Prior distributions or regularization methods (lasso penalties) are also accommodated. Missing values on dependent variables can be treated by full information maximum likelihood methods.
- A general (but experimental) Metropolis-Hastings sampler for Bayesian analysis based on MCMC is implemented in the `amh` function. Deterministic optimization of the posterior distribution (maximum posterior estimation or penalized maximum likelihood estimation) can be conducted with the `pmle` function which is based on `stats::optim`.

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Examples

```
## > library(LAM)
## ## LAM 0.0-4 (2017-03-03 16:53:46)
##
##      _ _      _ _ _ _ _      _ _      _ _
##  /\  \      /\  _ _  \  /\  "-./  \
##  \ \  _ _ _  \ \  _ _  \  \ \  \-./  \
##   \ _ _ _ _ \  \ _ _  \  \ _ _  \
##    \/_ _ _ _/   \/_/_/_/_/   \/_/_/   \/_/_/
##
```

amh

Bayesian Model Estimation with Adaptive Metropolis Hastings Sampling (amh) or Penalized Maximum Likelihood Estimation (pmle)

Description

The function `amh` conducts a Bayesian statistical analysis using the adaptive Metropolis-Hastings as the estimation procedure (Hoff, 2009). Only univariate prior distributions are allowed. Note that this function is intended just for experimental purpose, not to replace general purpose packages like **WinBUGS**, **JAGS**, **Stan** or **MHadaptive**.

The function `pmle` optimizes the penalized likelihood which means that the posterior is maximized and the maximum a posterior estimate is obtained. The optimization function `stats::optim` is used.

Usage

```
amh(data , nobs , pars , model , prior , proposal_sd , pars_lower = NULL ,
     pars_upper = NULL , derivedPars = NULL , n.iter = 5000 , n.burnin = 1000 ,
     n.sims=3000, acceptance_bounds = c(.45,.55), proposal_refresh = 50,
     print_iter = 50 )

pmle( data , nobs , pars , model , prior, pars_lower = NULL , pars_upper = NULL ,
      method = "L-BFGS-B" ,control=list() , verbose = TRUE , hessian = TRUE , ... )

## S3 method for class 'amh'
summary(object, digits = 3, file = NULL, ...)

## S3 method for class 'amh'
plot(x , conflevel=.95, digits=3 , lag.max= .1 ,
     col.smooth="red" , lwd.smooth=2 , col.split = "blue" , lwd.split = 2,
     lty.split=1, col.ci="orange", cex.summ=1, ask=FALSE , ... )

## S3 method for class 'amh'
coef(object, ...)

## S3 method for class 'amh'
logLik(object, ...)

## S3 method for class 'amh'
vcov(object, ...)

## S3 method for class 'amh'
confint(object , parm , level = .95, ... )

## S3 method for class 'pmle'
summary(object, digits = 3, file = NULL, ...)
```

```
## S3 method for class 'pmle'
coef(object, ...)

## S3 method for class 'pmle'
logLik(object, ...)

## S3 method for class 'pmle'
vcov(object, ...)

## S3 method for class 'pmle'
confint(object, parm, level = .95, ... )
```

Arguments

data	Object which contains data
nobs	Number of observations
pars	Named vector of initial values for parameters
model	Function defining the log-likelihood of the model
prior	List with prior distributions for the parameters to be sampled (see Examples). See sirt::prior_model_parse for more convenient specifications of the prior distributions.
proposal_sd	Vector with initial standard deviations for proposal distribution
pars_lower	Vector with lower bounds for parameters
pars_upper	Vector with upper bounds for parameters
derivedPars	Optional list containing derived parameters from sampled chain
n.iter	Number of iterations
n.burnin	Number of burn-in iterations
n.sims	Number of sampled iterations for parameters
acceptance_bounds	Bounds for acceptance probabilities of sampled parameters
proposal_refresh	Number of iterations for computation of adaptation of proposal standard deviation
print_iter	Display progress every print_iterth iteration
method	Optimization in stats::optim
control	Control parameters stats::optim
verbose	Logical indicating whether progress should be displayed.
hessian	Logical indicating whether the Hessian matrix should be computed
object	Object of class amh
digits	Number of digits used for rounding
file	File name

...	Further arguments to be passed
x	Object of class amh
confllevel	Confidence level
lag.max	Percentage of iterations used for calculation of autocorrelation function
col.smooth	Color moving average
lwd.smooth	Line thickness moving average
col.split	Color splitted chain
lwd.split	Line thickness splitted chain
lty.split	Line type splitted chain
col.ci	Color confidence interval
cex.summ	Point size summary
ask	Logical. If TRUE the user is asked for input, before a new figure is drawn.
parm	Optional vector of parameters.
level	Confidence level.

Value

List of class amh including entries

pars_chain	Data frame with sampled parameters
acceptance_parameters	Acceptance probabilities
amh_summary	Summary of parameters
coef	Coefficient obtained from marginal MAP estimation
pmle_pars	Object of parameters and posterior values corresponding to multivariate maximum of posterior distribution.
comp_estimators	Estimates for univariate MAP, multivariate MAP and mean estimator and corresponding posterior estimates.
ic	Information criteria
mcmcobj	Object of class mcmc for coda package
proposal_sd	Used proposal standard deviations
proposal_sd_history	History of proposal standard deviations during burn-in iterations
acceptance_rates_history	History of acceptance rates for all parameters during burn-in phase
...	More values

Author(s)

Alexander Robitzsch

References

Hoff, P. D. (2009). *A first course in Bayesian statistical methods*. New York: Springer.

See Also

See the Bayesian CRAN Task View for lot of information about alternative R packages.

[sirt::prior_model_parse](#)

Examples

```
## Not run:
#####
# EXAMPLE 1: Constrained multivariate normal distribution
#####

#--- simulate data
Sigma <- matrix( c(
  1 , .55 , .5 ,
  .55 , 1 , .45 ,
  .5 , .45 , 1 ) , nrow=3 , ncol=3 , byrow=TRUE )
mu <- c(0,1,1.2)
N <- 400
set.seed(9875)
dat <- MASS::mvrnorm( N , mu , Sigma )
colnames(dat) <- paste0("Y",1:3)
S <- stats::cov(dat)
M <- colMeans(dat)

#-- define maximum likelihood function for normal distribution
fit_m1 <- function( S , Sigma , M , mu , n , log=TRUE){
  Sigma1 <- solve(Sigma)
  p <- ncol(Sigma)
  det_Sigma <- det( Sigma )
  eps <- 1E-30
  if ( det_Sigma < eps ){
    det_Sigma <- eps
  }
  l1 <- - p * log( 2*pi ) - t( M - mu ) %*% Sigma1 %*% ( M - mu ) -
    log( det_Sigma ) - sum( diag( Sigma1 %*% S ) )
  l1 <- n/2 * l1
  if (! log){
    l1 <- exp(l1)
  }
  l1 <- l1[1,1]
  return(l1)
}
# This likelihood function can be directly accessed by the loglike_mvnorm function.

#--- define data input
data <- list( "S" = S , "M" = M , "n" = N )

#--- define list of prior distributions
```

```

prior <- list()
prior[["mu1"]] <- list( "dnorm" , list( x=NA , mean=0 , sd=1 ) )
prior[["mu2"]] <- list( "dnorm" , list( x=NA , mean=0 , sd=5 ) )
prior[["sig1"]] <- list( "dunif" , list( x=NA , 0 , 10 ) )
prior[["rho"]] <- list( "dunif" , list( x=NA , -1 , 1 ) )

*** alternatively, one can specify the prior as a string and uses
# the 'prior_model_parse' function
prior_model2 <- "
  mu1 ~ dnorm(x=NA, mean=0, sd=1)
  mu2 ~ dnorm(x=NA, mean=0, sd=5)
  sig1 ~ dunif(x=NA, 0,10)
  rho ~ dunif(x=NA,-1,1)
"

# convert string
prior2 <- sirt::prior_model_parse( prior_model2 )
prior2 # should be equal to prior

#--- define log likelihood function for model to be fitted
model <- function( pars , data ){
  # mean vector
  mu <- pars[ c("mu1", rep("mu2",2) ) ]
  # covariance matrix
  m1 <- matrix( pars["rho"] * pars["sig1"]^2 , 3 , 3 )
  diag(m1) <- rep( pars["sig1"]^2 , 3 )
  Sigma <- m1
  # evaluate log-likelihood
  ll <- fit_ml( S = data$S , Sigma = Sigma , M = data$M , mu = mu , n = data$n)
  return(ll)
}

#--- initial parameter values
pars <- c(1,2,2,0)
names(pars) <- c("mu1" , "mu2" , "sig1" , "rho")
#--- initial proposal distributions
proposal_sd <- c( .4 , .1 , .05 , .1 )
names(proposal_sd) <- names(pars)
#--- lower and upper bound for parameters
pars_lower <- c( -10 , -10 , .001 , -.999 )
pars_upper <- c( 10 , 10 , 1E100 , .999 )

#--- define list with derived parameters
derivedPars <- list( "var1" = ~ I( sig1^2 ) , "d1" = ~ I( ( mu2 - mu1 ) / sig1 ) )

*** start Metropolis-Hastings sampling
mod <- LAM::amh( data , nobs = data$n , pars=pars , model=model ,
  prior=prior , proposal_sd = proposal_sd ,
  n.iter = 1000 , n.burnin = 300 , derivedPars = derivedPars ,
  pars_lower = pars_lower , pars_upper = pars_upper )

# some S3 methods
summary(mod)
plot(mod, ask=TRUE)

```

```

coef(mod)
vcov(mod)
logLik(mod)

#--- compare Bayesian credibility intervals and HPD intervals
ci <- cbind( confint(mod) , coda::HPDinterval(mod$mcmcobj)[-1, ] )
ci
# interval lengths
cbind( ci[,2]-ci[,1] , ci[,4] - ci[,3] )

#--- plot update history of proposal standard deviations
graphics::matplot( x = rownames(mod$proposal_sd_history) ,
                   y = mod$proposal_sd_history , type="o" , pch=1:6)

#**** compare results with lavaan package
library(lavaan)
lavmodel <- "
  F~ 1*Y1 + 1*Y2 + 1*Y3
  F ~~ rho*F
  Y1 ~~ v1*Y1
  Y2 ~~ v1*Y2
  Y3 ~~ v1*Y3
  Y1 ~ mu1 * 1
  Y2 ~ mu2 * 1
  Y3 ~ mu2 * 1
  # total standard deviation
  sig1 := sqrt( rho + v1 )
"

# estimate model
mod2 <- lavaan::sem( data = as.data.frame(dat) , lavmodel )
summary(mod2)
logLik(mod2)

#*** compare results with penalized maximum likelihood estimation
mod3 <- LAM::pml( data=data , nobs= data$n , pars=pars , model = model , prior=prior ,
                 pars_lower = pars_lower , pars_upper = pars_upper , method = "L-BFGS-B" ,
                 control=list( trace=TRUE ) )

# model summaries
summary(res2)
confint(res2)
vcov(res2)

#--- lavaan with covariance and mean vector input
mod2a <- lavaan::sem( sample.cov = data$S , sample.mean = data$M , sample.nobs = data$n ,
                     model = lavmodel )

coef(mod2)
coef(mod2a)

#--- fit covariance and mean structure by fitting a transformed
#   covariance structure
#* create an expanded covariance matrix
p <- ncol(S)
S1 <- matrix( NA , nrow= p+1 , ncol=p+1 )

```



```

S1[1:p,1:p] <- S + outer( M , M )
S1[p+1,1:p] <- S1[1:p , p+1] <- M
S1[p+1,p+1] <- 1
vars <- c( colnames(S) , "MY" )
rownames(S1) <- colnames(S1) <- vars
#* lavaan model
lavmodel <- "
  # indicators
  F =~ 1*Y1 + 1*Y2 + 1*Y3
  # pseudo-indicator representing mean structure
  FM =~ 1*MY
  MY ~~ 0*MY
  FM ~~ 1*FM
  F ~~ 0*FM
  # mean structure
  FM =~ mu1*Y1 + mu2*Y2 + mu2*Y3
  # variance structure
  F ~~ rho*F
  Y1 ~~ v1*Y1
  Y2 ~~ v1*Y2
  Y3 ~~ v1*Y3
  sig1 := sqrt( rho + v1 )
"

# estimate model
mod2b <- lavaan::sem( sample.cov = S1 ,sample.nobs = data$n ,
  model = lavmodel )
summary(mod2b)
summary(mod2)

#####
# EXAMPLE 2: Estimation of a linear model with Box-Cox transformation of response
#####

*** simulate data with Box-Cox transformation
set.seed(875)
N <- 1000
b0 <- 1.5
b1 <- .3
sigma <- .5
lambda <- 0.3
# apply inverse Box-Cox transformation
# y1 = ( y^lambda - 1 ) / lambda
# -> y = ( lambda * y1 + 1 )^(1/lambda)
x <- stats::rnorm( N , mean=0 , sd = 1 )
y1 <- stats::rnorm( N , mean=b0 , sd = sigma ) + b1*x
# truncate at zero
eps <- .01
y1 <- ifelse( y1 < eps , eps , y1 )
y <- ( lambda * y1 + 1 ) ^ (1/lambda )

#-- display distributions of transformed and untransformed data
graphics::par(mfrow=c(1,2))

```

```

graphics::hist(yl, breaks=20)
graphics::hist(y, breaks=20)
graphics::par(mfrow=c(1,1))

#### define vector of parameters
pars <- c( 0 , 0 , 1 , -.2 )
names(pars) <- c("b0" , "b1" , "sigma" , "lambda" )
#### input data
data <- list( "y" = y , "x" = x )
#### define model with log-likelihood function
model <- function( pars , data ){
  sigma <- pars["sigma"]
  b0 <- pars["b0"]
  b1 <- pars["b1"]
  lambda <- pars["lambda"]
  if ( abs(lambda) < .01){ lambda <- .01 * sign(lambda) }
  y <- data$y
  x <- data$x
  n <- length(y)
  y_lambda <- ( y^lambda - 1 ) / lambda
  ll <- - n/2 * log(2*pi) - n * log( sigma ) -
    1/(2*sigma^2)* sum( (y_lambda - b0 - b1*x)^2 ) +
    ( lambda - 1 ) * sum( log( y ) )
  return(ll)
}
#-- test model function
model( pars , data )

#### define prior distributions
prior <- list()
prior[["b0"]] <- list( "dnorm" , list( x=NA , mean=0 , sd=10 ) )
prior[["b1"]] <- list( "dnorm" , list( x=NA , mean=0 , sd=10 ) )
prior[["sigma"]] <- list( "dunif" , list( x=NA , 0 , 10 ) )
prior[["lambda"]] <- list( "dunif" , list( x=NA , -2 , 2 ) )
#### define proposal SDs
proposal_sd <- c( .1 , .1 , .1 , .1 )
names(proposal_sd) <- names(pars)
#### define bounds for parameters
pars_lower <- c( -100 , -100 , .01 , -2 )
pars_upper <- c( 100 , 100 , 100 , 2 )

#### sampling routine
mod <- LAM::amh( data , nobs = N , pars , model , prior , proposal_sd ,
  n.iter = 10000 , n.burnin = 2000 , n.sims = 5000 ,
  pars_lower = pars_lower , pars_upper = pars_upper )
#-- S3 methods
summary(mod)
plot(mod , ask=TRUE )

#### estimating Box-Cox transformation in MASS package
library(MASS)
mod2 <- MASS::boxcox( stats::lm( y ~ x ) , lambda = seq(-1,2,length=100) )
mod2$x[ which.max( mod2$y ) ]

```

```

**** estimate Box-Cox parameter lambda with car package
library(car)
mod3 <- car::powerTransform( y ~ x )
summary(mod3)
# fit linear model with transformed response
mod3a <- stats::lm( car::bcPower( y, mod3$roundlam) ~ x )
summary(mod3a)

## End(Not run)

```

data.HT

Datasets from Heck and Thomas (2015)

Description

Selected datasets from Heck and Thomas (2015).

Usage

```
data(data.HT12)
```

Format

- The format of the dataset data.HT12 from Chapter 1 is:
 'data.frame': 120 obs. of 11 variables:
 \$ schcode: num 100 100 100 100 100 107 107 107 107 107 ...
 \$ read : num 682 644 651 710 673 593 660 640 646 634 ...
 \$ math : num 714 661 670 786 719 598 660 622 647 696 ...
 \$ lang : num 673 670 648 677 698 596 673 613 618 645 ...
 \$ ess : num -2.8 -2.8 -2.8 -2.8 -2.8 3.19 3.19 3.19 3.19 3.19 ...
 \$ cses : num -2.4 -2.4 -2.4 -2.4 -2.4 1.67 1.67 1.67 1.67 1.67 ...
 \$ female: num 0 0 0 0 0 1 1 1 0 ...
 \$ lowsesh: num 0 0 0 0 1 0 0 1 1 0 ...
 \$ lgsch : num 0 0 0 0 0 0 0 0 0 ...
 \$ age : num 135 140 135 151 138 138 140 141 144 146 ...
 \$ ncses : num 2.4 2.4 2.4 2.4 2.4 -1.67 -1.67 -1.67 -1.67 -1.67 ...

Source

<http://www2.hawaii.edu/~ltabata/HeckThomas2008/>, <https://www.routledge.com/An-Introduction-to-Multilevel-Modeling-Techniques-MLM-and-SEM-Approaches/Heck-Thomas/p/book/9781848725522>

References

Heck, R. H. & Thomas, S. L. (2015). *An introduction to multilevel modeling techniques*. Routledge, New York.

loglike_mvnorm

*Log-Likelihood Value of a Multivariate Normal Distribution***Description**

Computes log-likelihood value of a multivariate normal distribution given the empirical mean vector and the empirical covariance matrix as sufficient statistics.

Usage

```
loglike_mvnorm(M, S, mu, Sigma, n, log = TRUE, lambda = 0)
```

Arguments

M	Empirical mean vector
S	Empirical covariance matrix
mu	Population mean vector
Sigma	Population covariance matrix
n	Sample size
log	Optional logical indicating whether the logarithm of the likelihood should be calculated.
lambda	Regularization parameter of the covariance matrix (see Details).

Details

The population covariance matrix Σ is regularized if λ (lambda) is chosen larger than zero. Let Δ_{Σ} denote a diagonal matrix containing the diagonal entries of Σ . Then, a regularized matrix Σ^* is defined as $\Sigma^* = w\Sigma + (1 - w)\Delta_{\Sigma}$ with $w = n/(n + \lambda)$.

Value

Log-likelihood value

Author(s)

Alexander Robitzsch

Examples

```
#####
# EXAMPLE 1: Multivariate normal distribution
#####

library(MASS)

#--- simulate data
Sigma <- c( 1 , .55 , .5 , .55 , 1 , .5 , .5 , .5 , 1 )
```

```

Sigma <- matrix( Sigma , nrow=3 , ncol=3 )
mu <- c(0,1,1.2)
N <- 400
set.seed(9875)
dat <- MASS::mvrnorm( N , mu , Sigma )
colnames(dat) <- paste0("Y",1:3)
S <- cov(dat)
M <- colMeans(dat)

#--- evaluate likelihood
res1 <- LAM::loglike_mvnorm( M=M , S=S , mu=mu , Sigma=Sigma , n = N , lambda = 0 )
# compare log likelihood with somewhat regularized covariance matrix
res2 <- LAM::loglike_mvnorm( M=M , S=S , mu=mu , Sigma=Sigma , n = N , lambda = 1 )
print(res1)
print(res2)

```

mlnormal

*(Restricted) Maximum Likelihood Estimation with Prior Distributions
and Penalty Functions under Multivariate Normality*

Description

The `mlnormal` estimates statistical model for multivariate normally distributed outcomes with specified mean structure and covariance structure (see Details and Examples). Model classes include multilevel models, factor analysis, structural equation models, multilevel structural equation models, social relations model and perhaps more.

The estimation can be conducted under maximum likelihood, restricted maximum likelihood and maximum posterior estimation with prior distribution. Regularization (i.e. LASSO penalties) is also accommodated.

Usage

```

mlnormal(y, X, id, Z_list, Z_index, beta = NULL, theta, method = "ML", prior = NULL,
  lambda_beta = NULL, weights_beta = NULL, lambda_theta = NULL, weights_theta = NULL,
  beta_lower = NULL , beta_upper = NULL , theta_lower = NULL , theta_upper = NULL,
  maxit = 800, globconv = 1e-05, conv = 1e-06, verbose = TRUE, REML_shortcut = NULL,
  use_ginverse = FALSE, vcov = TRUE, variance_shortcut = TRUE, use_Rcpp = TRUE,
  level = 0.95, numdiff.parm = 1e-04, control_beta = NULL, control_theta = NULL)

## S3 method for class 'mlnormal'
summary(object, digits = 4, file = NULL, ...)

## S3 method for class 'mlnormal'
print(x, digits = 4, ...)

## S3 method for class 'mlnormal'
coef(object, ...)

```

```
## S3 method for class 'mlnormal'
logLik(object, ...)

## S3 method for class 'mlnormal'
vcov(object, ...)

## S3 method for class 'mlnormal'
confint(object , parm , level = .95, ... )
```

Arguments

y	Vector of outcomes
X	Matrix of covariates
id	Vector of identifiers (subjects or clusters, see Details)
Z_list	List of design matrices for covariance matrix (see Details)
Z_index	Array containing loadings of design matrices (see Details). The dimensions are units \times matrices \times parameters.
beta	Initial vector for β
theta	Initial vector for θ
method	Estimation method. Can be either "ML" or "REML".
prior	Prior distributions. Can be conveniently specified in a string which is processed by prior_model_parse . Only univariate prior distributions can be specified.
lambda_beta	Parameter λ_β for penalty function $P(\beta) = \lambda_\beta \sum_h w_{\beta h} \beta_h $
weights_beta	Parameter vector w_β for penalty function $P(\beta) = \lambda_\beta \sum_h w_{\beta h} \beta_h $
lambda_theta	Parameter λ_θ for penalty function $P(\theta) = \lambda_\theta \sum_h w_{\theta h} \theta_h $
weights_theta	Parameter vector w_θ for penalty function $P(\theta) = \lambda_\theta \sum_h w_{\theta h} \theta_h $
beta_lower	Vector containing lower bounds for β parameter
beta_upper	Vector containing upper bounds for β parameter
theta_lower	Vector containing lower bounds for θ parameter
theta_upper	Vector containing upper bounds for θ parameter
maxit	Maximum number of iterations
globconv	Convergence criterion deviance
conv	Maximum parameter change
verbose	Print progress?
REML_shortcut	Logical indicating whether computational shortcuts should be used for REML estimation
use_ginverse	Logical indicating whether a generalized inverse should be used
vcov	Logical indicating whether a covariance matrix of θ parameter estimates should be computed in case of REML (which is computationally demanding)
variance_shortcut	Logical indicating whether computational shortcuts for calculating covariance matrices should be used

use_Rcpp	Logical indicating whether the Rcpp package should be used
level	Confidence level
numdiff.parm	Numerical differentiation parameter
control_beta	List with control arguments for β estimation. The default is <code>list(maxiter=10, conv=1E-4, ridge = 1E-6)</code> .
control_theta	List with control arguments for θ estimation. The default is <code>list(maxiter=10, conv=1E-4, ridge = 1E-6)</code> .
object	Object of class <code>mlnormal</code>
digits	Number of digits used for rounding
file	File name
parm	Parameter to be selected for confint method
...	Further arguments to be passed
x	Object of class <code>mlnormal</code>

Details

The data consists of outcomes y_i and covariates X_i for unit i . The unit can be subjects, clusters (like schools) or the full outcome vector. It is assumed that y_i is normally distributed as $N(\mu_i, V_i)$ where the mean structure is modelled as

$$\mu_i = X_i\beta$$

and the covariance structure V_i depends on a parameter vector θ . More specifically, the covariance matrix V_i is modelled as a sum of functions of the parameter θ and known design matrices Z_{im} for unit i ($m = 1, \dots, M$). The model is

$$V_i = \sum_{m=1}^M \gamma_{im} Z_{im} \quad \text{with} \quad \gamma_{im} = \prod_{h=1}^H \theta_h^{q_{imh}}$$

where q_{imh} are non-negative known integers specified in `Z_index` and Z_{im} are design matrices specified in `Z_list`.

The estimation follows Fisher scoring (Jiang, 2007; for applications see also Longford, 1987; Lee, 1990; Gill & Swartz, 2001) and the regularization approach is as described in Lin, Pang and Jiang (2013) (see also Krishnapuram, Carin, Figueiredo, & Hartemink, 2005).

Value

List with entries

theta	Estimated θ parameter
beta	Estimated β parameter
theta_summary	Summary of θ parameters
beta_summary	Summary of β parameters
coef	Estimated parameters
vcov	Covariance matrix of estimated parameters

ic	Information criteria
V_list	List with fitted covariance matrices V_i
Vl_list	List with inverses of fitted covariance matrices V_i
prior_args	Some arguments in case of prior distributions
...	More values

Author(s)

Alexander Robitzsch

References

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- Longford, N. T. (1987). A fast scoring algorithm for maximum likelihood estimation in unbalanced mixed models with nested random effects. *Biometrika*, **74**, 817-827.

See Also

See **lavaan**, **sem**, **lava**, **OpenMx** or **nlsem** packages for estimation of (single level) structural equation models.

See the **regsem** and **lsf** packages for regularized structural equation models.

See **lme4** or **nlme** package for estimation of multilevel models.

See the **lmmlasso** and **glmmLasso** packages for regularized mixed effects models.

See **OpenMx** and **xxM** packages (<http://xxm.times.uh.edu/>) for estimation of multilevel structural equation models.

Examples

```
## Not run:
#####
# EXAMPLE 1: Two-level random intercept model
#####

#-----
# Simulate data
#-----
```



```

set.seed(976)
G <- 150 ; rg <- c(10,20) # 150 groups with group sizes ranging from 10 to 20
## simulate group sizes
ng <- round( stats::runif( G , min = rg[1] , max = rg[2] ) )
idcluster <- rep(1:G , ng )
## simulate covariate
iccx <- .3
x <- rep( stats::rnorm( G , sd = sqrt( iccx ) , ng ) +
          stats::rnorm( sum(ng) , sd = sqrt( 1 - iccx ) )
## simulate outcome
b0 <- 1.5 ; b1 <- .4 ; iccy <- .2
y <- b0 + b1*x + rep( stats::rnorm( G , sd = sqrt( iccy ) , ng ) +
                      stats::rnorm( sum(ng) , sd = sqrt( 1 - iccy ) )

#-----
#--- arrange input for mlnormal function

id <- idcluster          # cluster is identifier
X <- cbind( 1 , x )      # matrix of covariates
N <- length(id)          # number of units (clusters), which is G

MD <- max(ng)            # maximum number of persons in a group
NP <- 2                  # number of covariance parameters theta

## list of design matrix for covariance matrix
# In the case of the random intercept model, the covariance structure is
#  $\tau^2 * J + \sigma^2 * I$ , where J is a matrix of ones and I is the
# identity matrix
Z <- as.list(1:G)
for (gg in 1:G){
  Ngg <- ng[gg]
  Z[[gg]] <- as.list( 1:2 )
  Z[[gg]][[1]] <- matrix( 1 , nrow=Ngg , ncol=Ngg ) # level 2 variance
  Z[[gg]][[2]] <- diag(1,Ngg)                      # level 1 variance
}
Z_list <- Z
## parameter list containing the powers of parameters
Z_index <- array( 0 , dim=c(G,2,2) )
Z_index[ 1:G , 1 , 1] <- Z_index[ 1:G , 2 , 2] <- 1

##** starting values and parameter names
beta <- c( 1 , 0 )
names(beta) <- c("int" , "x")
theta <- c( .05 , 1 )
names(theta) <- c("tau2" , "sig2" )

##** create dataset for lme4 for comparison
dat <- data.frame(y=y , x = x , id = id )

#-----
# Model 1: Maximum likelihood estimation
#-----

```

```

*** mlnormal function
mod1a <- LAM::mlnormal( y =y, X=X , id=id , Z_list=Z_list , Z_index =Z_index,
                      beta = beta , theta = theta, method= "ML" )
summary(mod1a)

# lme4::lmer function
library(lme4)
mod1b <- lme4::lmer( y ~ x + (1 | id ) , data = dat , REML = FALSE )
summary(mod1b)

#-----
# Model 2: Restricted maximum likelihood estimation
#-----

*** mlnormal function
mod2a <- LAM::mlnormal( y =y, X=X , id=id , Z_list=Z_list , Z_index =Z_index,
                      beta = beta , theta = theta, method= "REML" )
summary(mod2a)

# lme4::lmer function
mod2b <- lme4::lmer( y ~ x + (1 | id ) , data = dat , REML = TRUE )
summary(mod2b)

#-----
# Model 3: Estimation of standard deviation instead of variances
#-----

# The model is now parametrized in standard deviations
# Variances are then modelled as tau^2 and sigma^2, respectively.
Z_index2 <- 2*Z_index      # change loading matrix

# estimate model
mod3 <- LAM::mlnormal( y =y, X=X , id=id , Z_list=Z_list , Z_index =Z_index2,
                      beta = beta , theta = theta )
summary(mod3)

#-----
# Model 4: Maximum posterior estimation
#-----

# specify prior distributions for parameters
prior <- "
  tau2 ~ dgamma(NA , 2 , .5 )
  sig2 ~ dinvgamma(NA , .1 , .1 )
  x ~ dnorm( NA , .2 , 1000 )
"

# estimate model in mlnormal
mod4 <- LAM::mlnormal( y =y, X=X , id=id , Z_list=Z_list , Z_index =Z_index,
                      beta = beta, theta = theta, method= "REML", prior = prior, vcov = FALSE )
summary(mod4)

#-----

```

```

# Model 5: Estimation with regularization on beta and theta parameters
#-----

*** penalty on theta parameter
lambda_theta <- 10
weights_theta <- 1 + 0 * theta
*** penalty on beta parameter
lambda_beta <- 3
weights_beta <- c( 0 , 1.8 )

# estimate model
mod5 <- LAM::mlnormal( y=y, X=X, id=id , Z_list=Z_list , Z_index=Z_index,
                      beta=beta, theta=theta, method= "ML" , maxit=maxit ,
                      lambda_theta=lambda_theta, weights_theta=weights_theta ,
                      lambda_beta=lambda_beta, weights_beta=weights_beta )
summary(mod5)

#####
# EXAMPLE 2: Latent covariate model, two-level regression
#####

#  $Y_b = \beta_0 + \beta_b X_b + e_b$  (between level) and
#  $Y_w = \beta_w X_w + e_w$  (within level)

#-----
# Simulate data from latent covariate model
#-----

set.seed(865)
# regression parameters
beta_0 <- 1 ; beta_b <- .7 ; beta_w <- .3
G <- 200      # number of groups
n <- 15       # group size
iccx <- .2    # intra class correlation x
iccy <- .35   # (conditional) intra class correlation y
# simulate latent variables
xb <- stats::rnorm(G , sd = sqrt( iccx ) )
yb <- beta_0 + beta_b * xb + stats::rnorm(G , sd = sqrt( iccy ) )
xw <- stats::rnorm(G*n , sd = sqrt( 1-iccx ) )
yw <- beta_w * xw + stats::rnorm(G*n , sd = sqrt( 1-iccy ) )
group <- rep( 1:G , each = n )
x <- xw + xb[ group ]
y <- yw + yb[ group ]
# test results on true data
lm( yb ~ xb )
lm( yw ~ xw )

# create vector of outcomes in the form
# ( y_11 , x_11 , y_21 , x_21 , ... )
dat <- cbind( y , x )
dat
Y <- as.vector( t(dat) )    # outcome vector
ny <- length(Y)

```

```

X <- matrix( 0 , nrow=ny , ncol=2 )
X[ seq(1,ny,2) , 1 ] <- 1 # design vector for mean y
X[ seq(2,ny,2) , 2 ] <- 1 # design vector for mean x
id <- rep( group , each = 2 )

#-----
# Model 1: Linear regression ignoring multilevel structure
#-----

# y = beta_0 + beta_t * x + e
# Var(y) = beta_t^2 * var_x + var_e
# Cov(y,x) = beta_t * var_x
# Var(x) = var_x

*** initial parameter values
theta <- c( 0 , 1 , .5 )
names(theta) <- c( "beta_t" , "var_x" , "var_e")
beta <- c(0,0)
names(beta) <- c("mu_y","mu_x")

# The unit i is a cluster in this example.

#--- define design matrices | list Z_list
Hlist <- list( matrix( c(1,0,0,0) , 2 , 2 ) , # var(y)
              matrix( c(1,0,0,0) , 2 , 2 ) , # var(y) (two terms)
              matrix( c(0,1,1,0) , 2 , 2 ) , # cov(x,y)
              matrix( c(0,0,0,1) , 2 , 2 ) ) # var(x)

U0 <- matrix( 0 , nrow=2*n , ncol=2*n )
Ulist <- list( U0 , U0 , U0 , U0 )
M <- length(Hlist)
for (mm in 1:M){ # mm <- 1
  for (nn in 1:n){ # nn <- 0
    Ulist[[ mm ]][ 2*(nn-1) + 1:2 , 2*(nn-1) + 1:2 ] <- Hlist[[ mm ]]
  }
}
Z_list <- as.list(1:G)
for (gg in 1:G){
  Z_list[[gg]] <- Ulist
}

#--- define index vectors
Z_index <- array( 0 , dim=c(G , 4 , 3 ) )
K0 <- matrix( 0 , nrow=4 , ncol=3 )
colnames(K0) <- names(theta)
# Var(y) = beta_t^2 * var_x + var_e (matrices withn indices 1 and 2)
K0[ 1 , c("beta_t","var_x") ] <- c(2,1) # beta_t^2 * var_x
K0[ 2 , c("var_e") ] <- c(1) # var_e
# Cov(y,x) = beta_t * var_x
K0[ 3 , c("beta_t","var_x")] <- c(1,1)
# Var(x) = var_x
K0[ 4 , c("var_x") ] <- c(1)
for (gg in 1:G){

```

```

      Z_index[gg,,] <- K0
    }

    **** estimate model with mlnormal
    mod1a <- LAM::mlnormal( y=Y, X=X , id=id , Z_list=Z_list , Z_index=Z_index,
                          beta=beta, theta=theta, method="REML" , vcov=FALSE )
    summary(mod1a)

    **** estimate linear regression with stats::lm
    mod1b <- stats::lm( y ~ x )
    summary(mod1b)

    #-----
    # Model 2: Latent covariate model
    #-----

    *** initial parameters
    theta <- c( 0.12 , .6 , .5 , 0 , .2 , .2 )
    names(theta) <- c( "beta_w" , "var_xw" , "var_ew" ,
                      "beta_b" , "var_xb" , "var_eb" )

    #--- define design matrices | list Z_list
    Hlist <- list( matrix( c(1,0,0,0) , 2 , 2 ) , # var(y)
                  matrix( c(1,0,0,0) , 2 , 2 ) , # var(y) (two terms)
                  matrix( c(0,1,1,0) , 2 , 2 ) , # cov(x,y)
                  matrix( c(0,0,0,1) , 2 , 2 ) ) # var(x)
    U0 <- matrix( 0 , nrow=2*n , ncol=2*n )
    Ulist <- list( U0 , U0 , U0 , U0 , # within structure
                  U0 , U0 , U0 , U0 ) # between structure
    M <- length(Hlist)
    **** within structure
    design_within <- diag(n) # design matrix within structure
    for (mm in 1:M){ # mm <- 1
      Ulist[[ mm ]] <- base::kronecker( design_within , Hlist[[mm]] )
    }
    **** between structure
    design_between <- matrix(1, nrow=n , ncol=n)
      # matrix of ones corresponding to group size
    for (mm in 1:M){ # mm <- 1
      Ulist[[ mm + M ]] <- base::kronecker( design_between , Hlist[[ mm ]] )
    }
    Z_list <- as.list(1:G)
    for (gg in 1:G){
      Z_list[[gg]] <- Ulist
    }

    #--- define index vectors Z_index
    Z_index <- array( 0 , dim=c(G , 8 , 6 ) )
    K0 <- matrix( 0 , nrow=8 , ncol=6 )
    colnames(K0) <- names(theta)
    # Var(y) = beta^2 * var_x + var_e (matrices withn indices 1 and 2)
    K0[ 1 , c("beta_w","var_xw") ] <- c(2,1) # beta_t^2 * var_x
    K0[ 2 , c("var_ew") ] <- c(1) # var_e

```

```

K0[ 5 , c("beta_b","var_xb") ] <- c(2,1) # beta_t^2 * var_x
K0[ 6 , c("var_eb") ] <- c(1) # var_e
# Cov(y,x) = beta * var_x
K0[ 3 , c("beta_w","var_xw")] <- c(1,1)
K0[ 7 , c("beta_b","var_xb")] <- c(1,1)
# Var(x) = var_x
K0[ 4 , c("var_xw") ] <- c(1)
K0[ 8 , c("var_xb") ] <- c(1)
for (gg in 1:G){
  Z_index[gg,,] <- K0
}

#--- estimate model with mlnormal
mod2a <- LAM::mlnormal( y=Y, X=X , id=id , Z_list=Z_list , Z_index=Z_index,
  beta=beta , theta=theta, method="ML" )
summary(mod2a)

#####
# EXAMPLE 3: Simple linear regression, single level
#####

#-----
# Simulate data
#-----

set.seed(875)
N <- 300
x <- stats::rnorm( N , sd = 1.3 )
y <- .4 + .7 * x + stats::rnorm( N , sd = .5 )
dat <- data.frame( x , y )

#-----
# Model 1: Linear regression modelled with residual covariance structure
#-----

# matrix of predictros
X <- cbind( 1 , x )
# list with design matrices
Z <- as.list(1:N)
for (nn in 1:N){
  Z[[nn]] <- as.list( 1 )
  Z[[nn]][[1]] <- matrix( 1 , nrow=1, ncol=1 ) # residual variance
}
#* loading matrix
Z_index <- array( 0 , dim=c(N,1,1) )
Z_index[ 1:N , 1 , 1] <- 2 # parametrize residual standard deviation
##* starting values and parameter names
beta <- c( 0 , 0 )
names(beta) <- c("int" , "x")
theta <- c(1)
names(theta) <- c("sig2" )
# id vector
id <- 1:N

```

```

*** mlnormal function
mod1a <- LAM::mlnormal( y=y, X=X , id=id , Z_list=Z , Z_index =Z_index,
                      beta = beta , theta = theta, method= "ML" )
summary(mod1a)

# estimate linear regression with stats::lm
mod1b <- stats::lm( y ~ x )
summary(mod1b)

#-----
# Model 2: Linear regression modelled with bivariate covariance structure
#-----

*** define design matrix referring to mean structure
X <- matrix( 0 , nrow=2*N , ncol=2 )
X[ seq(1,2*N,2) , 1 ] <- X[ seq(2,2*N,2) , 2 ] <- 1

*** create outcome vector
y1 <- dat[ cbind( rep(1:N, each=2) , rep(1:2, N ) ) ]
*** list with design matrices
Z <- as.list(1:N)
Z0 <- 0*matrix( 0 , nrow=2,ncol=2)
ZXY <- ZY <- ZX <- Z0
# design matrix Var(X)
ZX[1,1] <- 1
# design matrix Var(Y)
ZY[2,2] <- 1
# design matrix covariance
ZXY[1,2] <- ZXY[2,1] <- 1
# Var(X) = sigx^2
# Cov(X,Y) = beta * sigx^2
# Var(Y) = beta^2 * sigx^2 + sige^2
Z_list0 <- list( ZY , ZY , ZXY , ZX )
for (nn in 1:N){
  Z[[nn]] <- Z_list0
}
#* parameter list containing the powers of parameters
theta <- c(1,0.3,1)
names(theta) <- c("sigx", "beta" , "sige" )
Z_index <- array( 0 , dim=c(N,4,3) )
for (nn in 1:N){
  # Var(X)
  Z_index[nn, 4 , ] <- c(2,0,0)
  # Cov(X,Y)
  Z_index[nn, 3, ] <- c(2,1,0)
  # Var(Y)
  Z_index[nn,1,] <- c(2,2,0)
  Z_index[nn,2,] <- c(0,0,2)
}
#* starting values and parameter names
beta <- c( 0 , 0 )
names(beta) <- c("Mx" , "My")

```

```

# id vector
id <- rep( 1:N , each=2 )

*** mlnormal function
mod2a <- LAM::mlnormal( y =y1, X=X , id=id , Z_list=Z , Z_index =Z_index,
                      beta = beta , theta = theta, method= "ML" )
summary(mod2a)

#-----
# Model 3: Bivariate normal distribution in (sigma_X, sigma_Y, sigma_XY) parameters
#-----

# list with design matrices
Z <- as.list(1:N)
Z0 <- 0*matrix( 0 , nrow=2,ncol=2)
ZXY <- ZY <- ZX <- Z0
# design matrix Var(X)
ZX[1,1] <- 1
# design matrix Var(Y)
ZY[2,2] <- 1
# design matrix covariance
ZXY[1,2] <- ZXY[2,1] <- 1
Z_list0 <- list( ZX , ZY , ZXY )
for (nn in 1:N){
  Z[[nn]] <- Z_list0
}

#* parameter list
theta <- c(1,1,.3)
names(theta) <- c("sigx", "sigy" , "sigxy" )
Z_index <- array( 0 , dim=c(N,3,3) )
for (nn in 1:N){
  # Var(X)
  Z_index[nn, 1 , ] <- c(2,0,0)
  # Var(Y)
  Z_index[nn, 2 , ] <- c(0,2,0)
  # Cov(X,Y)
  Z_index[nn, 3, ] <- c(0,0,1)
}

*** starting values and parameter names
beta <- c( 0 , 0 )
names(beta) <- c("Mx" , "My")

*** mlnormal function
mod3a <- LAM::mlnormal( y =y1, X=X , id=id , Z_list=Z , Z_index =Z_index,
                      beta = beta , theta = theta, method= "ML" )
summary(mod3a)

#-----
# Model 4: Bivariate normal distribution in parameters of Cholesky decomposition
#-----

```



```

# list with design matrices
Z <- as.list(1:N)
Z0 <- 0*matrix( 0 , nrow=2,ncol=2)
ZXY <- ZY <- ZX <- Z0
# design matrix Var(X)
ZX[1,1] <- 1
# design matrix Var(Y)
ZY[2,2] <- 1
# design matrix covariance
ZXY[1,2] <- ZXY[2,1] <- 1
Z_list0 <- list( ZX , ZXY , ZY , ZY )
for (nn in 1:N){
  Z[[nn]] <- Z_list0
}

## parameter list containing the powers of parameters
theta <- c(1,0.3,1)
names(theta) <- c("L11", "L21" , "L22" )
Z_index <- array( 0 , dim=c(N,4,3) )
for (nn in 1:N){
  Z_index[nn,1,] <- c(2,0,0)
  Z_index[nn,2,] <- c(1,1,0)
  Z_index[nn,3,] <- c(0,2,0)
  Z_index[nn,4,] <- c(0,0,2)
}
##* starting values and parameter names
beta <- c( 0 , 0 )
names(beta) <- c("Mx" , "My")
# id vector
id <- rep( 1:N , each=2 )
##* mlnormal function
mod4a <- LAM::mlnormal( y =y1, X=X , id=id , Z_list=Z , Z_index =Z_index,
  beta = beta , theta = theta, method= "ML" )
# parameter with lower diagonal entries of Cholesky matrix
mod4a$theta
# fill-in parameters for Cholesky matrix
L <- matrix(0,2,2)
L[ ! upper.tri(L) ] <- mod4a$theta
##* reconstruct covariance matrix
L
stats::cov.wt(dat, method="ML")$cov

## End(Not run)

```

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