Package ‘exact2x2’

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Type Package

Title Exact Tests and Confidence Intervals for 2x2 Tables

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Depends R (>= 2.10), stats (>= 3.1.1), exactci, ssanv

Description Calculates conditional exact tests (Fisher’s exact test, Blaker’s exact test, or exact McNemar’s test) and unconditional exact tests (including score-based tests on differences in proportions, ratios of proportions, and odds ratios, and Boshcloo’s test) with appropriate matching confidence intervals, and provides power and sample size calculations. Gives melded confidence intervals for the binomial case. Gives boundary-optimized rejection region test (Gabriel, et al, 2018, <DOI:10.1002/sim.7579>), an unconditional exact test for the situation where the controls are all expected to fail.

License GPL-3

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R topics documented:

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Description

There are 7 main functions in the package. The `exact2x2` function calculates the exact conditional tests with matching confidence intervals as detailed in Fay (2010a <DOI:10.1093/biostatistics/kxp050>, 2010b). The functions `ss2x2` and `power2x2` calculate the sample size and power related to the tests of `exact2x2`. The `uncondExact2x2` and `boschloo` functions calculate unconditional exact tests. The `binommeldNtest` function calculates melded confidence intervals for two sample binomial inferences (see Fay, Proschan, and Brittain, 2015 <DOI:10.1111/biom.12231>). Finally, the `bortTest` function calculates the boundary optimized rejection region test that creates unconditional exact tests that have power optimized when group 1 is expected to have 100 percent failure. For example, in vaccine challenge studies where the control group are all expected to get infected (see Gabriel, et al, 2018 <DOI:10.1002/sim.7579>, the letter about that paper by Martin Andres <DOI:10.1002/sim.7630>, and the response <DOI:10.1002/sim.7684>).

Details

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Author(s)

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References


**binomMeld.test**

*Melded Binomial Confidence Intervals and Tests*

**Description**

Creates tests to compare two binomials, giving confidence intervals for either the difference in proportions, the rate ratio, or the odds ratio. The 95 percent confidence intervals have been shown to guarantee nominal coverage by extensive numerical calculations. It has been theoretically proven that the p-values from the one-sided tests on the null hypothesis of equality match Fisher’s exact p-values.

**Usage**

```r
binomMeld.test(x1, n1, x2, n2, nullparm = NULL, 
parmtype = c("difference", "oddsratio", "ratio"),
conf.level = 0.95, conf.int=TRUE,
alternative = c("two.sided", "less", "greater"),
midp=FALSE, nmc=0, eps=10^-8)
```

**Arguments**

- `x1`: number of events in group 1
- `n1`: sample size in group 1
- `x2`: number of events in group 2
- `n2`: sample size in group 2
- `nullparm`: value of the parameter of interest at null, default of NULL gives 0 for parmtype='difference' and 1 for parmtype='ratio' or 'oddsratio'
- `parmtype`: type of parameter of interest, one of "difference", "ratio" or "oddsratio" (see details)
- `conf.level`: confidence level
- `conf.int`: logical, calculate confidence intervals?
- `alternative`: alternative hypothesis, one of "two.sided", "less", or "greater" (see details)
- `midp`: logical, do mid-p version of p-value and confidence intervals?
- `nmc`: integer, number of Monte Carlo replications for p-value and CI calculations, 0 (default) means calculate by numeric integration instead
- `eps`: small number used to adjust numeric integration (see note)
Details
Assume $X_1 \sim \text{Binomial}(n_1, p_1)$ and $X_2 \sim \text{Binomial}(n_2, p_2)$. We want to test hypotheses on a function of $p_1$ and $p_2$. The functions are given by \texttt{parmtype}: difference tests $p_2-p_1$, ratio tests $p_2/p_1$, and odds ratio tests $(p_2(1-p_1))/(p_1(1-p_2))$. Let $g(p_1,p_2)$ be one of the three functions. So when alternative is "less" we test $H_0: g(p_1,p_2) \geq \text{nullparm}$ vs. $H_1: g(p_1,p_2) < \text{nullparm}$.

For details when \texttt{midp}\texttt{=}FALSE see Fay, Proschan, and Brittain (2015).

When \texttt{midp}\texttt{=}TRUE, the method performs the mid-p version on the p-value and the associated confidence intervals. This means that we replace the confidence distribution random variables in the p-value and CI calculations with a random variable that is a mixture of the lower and upper CD random variables. For example, if $W_1L$ and $W_1U$ are the lower and upper confidence distribution random variables for group 1, then we replace those values in all calculations with $W_1\text{midp} = U_1 \times W_1L + (1-U_1) \times W_1U$, where $U_1$ is a Bernoulli with parameter 0.5. For a discussion of mid-p p-values and the associated confidence intervals in a closely related context, see the vignette on mid-p values or Fay and Brittain (2016, especially the Appendix).

Value
An object of class 'htest'. A list with elements

- \texttt{statistic} proportion of events in group 1
- \texttt{parameter} proportion of events in group 2
- \texttt{p.value} p-value
- \texttt{conf.int} confidence interval
- \texttt{estimate} estimate of $g(p_1,p_2)$ by plugging in sample proportions, i.e., unconditional MLE
- \texttt{null.value} value of $g(p_1,p_2)$ under null
- \texttt{alternative} type of alternative hypothesis
- \texttt{method} description of test
- \texttt{data.name} character explicit description of data

Note
For numeric integration, the integrate function may have problems if nearly all of the integrand values are about 0 within the range of integration. Because of this, we use the eps value to make sure we integrate over ranges in which the integrand is nontrivially greater than 0. We restrict the range then add eps back to the p-value so that if the integrate function works perfectly, then the p-values would be very slightly conservative (for very small eps). There is no need to adjust the eps value. See code for detailed description of how eps is used in the calculation before changing it from the default.

An alternative method of calculation is to use Monte Carlo simulation (option with \texttt{nmc}>0). This provides a check of the numeric integration. There is no need to do Monte Carlo simulations for routine use. Please inform the package maintainer if the p-values or confidence intervals are substantially different when \texttt{nmc}=0 and \texttt{nmc}=10^7.

Author(s)
Michael P. Fay
**borrTest**

**Boundary-Optimized Rejection Region Test**

**Description**

An unconditional exact test for the two-sample binomial problem when it is expected that theta1 (probability of an event in group 1) will be close to 1. Used for test versus control when all controls are expected to fail.

**Usage**

```r
borrTest(x1, n1, x2, n2, tuningParm = 0.025,
  parmtype = c("ratio", "difference", "oddsratio"),
  nullparm = NULL, alternative = c("less", "greater", "two.sided"),
  conf.int = TRUE, conf.level = 0.975,
  controlUC = ucControl(), controlborr = borrControl(), ...)

borrPvals(n1,n2, tuningParm=0.025,
  parmtype = c("ratio", "difference","oddsratio"),
  nullparm = NULL, alternative = c("less", "greater","two.sided"),
  conf.int = TRUE, conf.level = 0.975,
  controlUC=ucControl(), controlborr=borrControl(),...)

borrOrdering(n1,n2,tuningParm = .025,
  controlborr=borrControl())
```

**References**


**Examples**

```r
# Note the p-value for all tests of equality
# (Null Hypothesis: true prop 1=true prop 2)
# are the same, and equal to the
# Fisher's exact (central) p-value
binomMeld.test(3,5,1,8,parmtype="difference")
binomMeld.test(3,5,1,8,parmtype="ratio")
# note that binomMeld.test gives the unconditional MLE
# for the odds ratio, while fisher.test and exact2x2
# gives the conditional MLE for the odds ratio
# (also fisher.test gives the odds ratio defined as
# the inverse of how it is defined in binomMeld.test)
binomMeld.test(3,5,1,8,parmtype="oddsratio")
extact2x2(matrix(c(1,8-1,3,5-3),2,2),tmethod="central")
```
Arguments

x1 number of events in group 1
n1 sample size in group 1
x2 number of events in group 2
n2 sample size in group 2

tuningParm tuning parameter, default is 0.025 and designs BORR tests with maximum power for one-sided 0.025 tests

parmtype parameter type, either 'ratio' for theta2/theta1, 'difference' for theta2-theta1, or 'oddsratio' for theta2*(1-theta1)/(theta1*(1-theta2)).

nullparm null parameter value, default=NULL gives parameter value for theta1=theta2 (e.g., 1 for 'ratio' or 0 for 'difference').

alternative alternative hypothesis, BORR tests are designed for alternative='less' (see Note for other alternatives)

conf.int logical, should confidence interval be calculated?
conf.level confidence level, default is 0.975 (see note)

controlUC a list of control parameters to define algorithms in the call to uncondExact2x2, see ucControl

controlburr a list of control parameters to define algorithms, see borrControl

... extra arguments passed (for future development)

Details

The boundary-optimized rejection region test is designed to test the one-sided alternative that theta2 < theta1, where X1 is binomial(n1,theta1), and X2 is binomial(n2,theta2). The test is designed to be optimal when theta1 is very close to 1. For example, in a vaccine malaria challenge study where we expect all n1 individuals that got the control vaccine to have the event (get malaria when challenged with malaria). For details see Gabriel et al (2018).

The function `borrTest` tests the results of one study, and returns an `htest` object. The function `borrPvals` calculates the p-values for every possible result of a study. The function `borrOrdering` orders every possible result of the study. See `borrOrderingInternal` for calculation details.

Value

The function `borrPvals` returns a (n1+1) by (n2+1) matrix of p-values for all possible x1 and x2 values. The function `borrOrdering` returns a matrix with the rank of all possible x1 and x2 values. The function `borrTest` returns a list of class `htest` with elements:

- statistic proportion in sample 1
- parameter proportion in sample 2
- p.value p-value from test
- conf.int confidence interval on parameter given by parmtype
- estimate MLE estimate of parameter given by parmtype
- null.value null hypothesis value of parameter given by parmtype
alternative alternative hypothesis
method description of test
data.name description of data

Note

The tests are designed to have good power for the one-sided test that $H_0: \theta_2 \geq \theta_1$, with alternative $H_1: \theta_2 < \theta_1$ at significance level equal to $tuningParm$. Since the default $tuningParm$ is 0.025, the default confidence level is 0.975 so that the confidence intervals will be compatible with the test where the one-sided p-values reject at level 0.025.

Sometimes you may want two-sided confidence intervals on the parameter of interest. If you ask for a two-sided alternative, then the confidence interval and the resulting p-value will be two-sided as well. The default is a 'central' interval, so the two-sided p-value should be twice the minimum of the one-sided p-values. Further, with a $conf.level$ of 0.95 for the two-sided alternative, the error on each side will be bounded by 0.025.

Author(s)

Martha Nason, Erin Gabriel, Michael P. Fay

References


Examples

```r
## Not run: borrTest(4,4,1,4)
# Note Figure 2 in Gabriel et al is incorrect. The correct value
# is in the response letter, and given by
# borrOrdering(4,4,tuningParm=0.025)$rankMat
```

boschloo

Boschloo’s test for 2x2 Tables

Description

Boschloo’s test is an exact unconditional test for 2x2 tables based on ordering the sample space by Fisher’s exact p-values. This function generalizes that test in several ways (see details).
Usage

boschloo(x1, n1, x2, n2, alternative = c("two.sided", "less", "greater"),
or = NULL, conf.int = FALSE, conf.level = 0.95, midp = FALSE,
tsmethod = c("central", "minlike"), control=ucControl())

Arguments

x1 number of events in group 1
n1 sample size in group 1
x2 number of events in group 2
n2 sample size in group 2
alternative alternative hypothesis, one of "two.sided", "less", or "greater", default is "two.sided" (see details)
or odds ratio under the null hypothesis
conf.int logical, calculate confidence interval?
conf.level confidence level
midp logical. Use mid-p-value method?
tsmethod two-sided method, either "central" or "minlike" (see details)
control list of algorithm parameters, see ucControl

Details

The traditional Boschloo (1970) test is to use Fisher’s exact p-values (under the null that p1=p2) to order the sample space and to use that ordering to perform an unconditional exact test. Here we generalize this to test for different null hypothesis values (other than odds ratios of 1).

For the two-sided alternatives, the traditional method uses tsmethod='minlike' (for example, in the Exact R package) but our default is tsmethod='central'. The one-sided tests use ordering by the appropriate p-value (or 1 minus the p-value for alternative='greater' so that the ordering function follows our convention for user supplied ordering functions, see method='user' option in uncondExact2x2).

The option midp orders the sample space by the mid-p value associated with Fisher’s exact test, and additionally gives mid-p values. This means that unlike the midp=FALSE case, when midp=TRUE the test is not exact (i.e., guaranteed to bound the type I error rate at the nominal level), but has type I error rates that are on average (over the possible null parameter values) closer to the nominal level.

If you want to order by the mid-p values from Fisher’s exact test but get an exact test, use the method="FisherAdj" with midp=FALSE in uncondExact2x2.

The boschloo function only gives confidence intervals for the odds ratio, for getting confidence intervals closely related to Boschloo p-values (but not exactly matching Boschloo p-values) for the difference or ratio, use uncondExact2x2 with method="FisherAdj".

exact2x2

Value

a list of class 'htest' with elements:

- statistic: proportion in sample 1
- parameter: proportion in sample 2
- p.value: p-value from test
- conf.int: confidence interval on odds ratio
- estimate: odds ratio estimate
- null.value: null hypothesis value of odds ratio
- alternative: alternative hypothesis
- method: description of test
- data.name: description of data

References


See Also

exact.test in package Exact for Boschloo test p-value computation. Also see method"FisherAdj" in uncondExact2x2 for a closely related test.

Examples

# defaults to the central two-sided version
boschloo(1,5,6,7)
boschloo(1,5,6,7,alternative="greater")
## traditional two-sided Boschloo test (not central!)
boschloo(1,5,6,7, tsmethod="minlike")

---

**exact2x2**  
*Exact Conditional Tests for 2 by 2 Tables of Count Data*

Description

Performs exact conditional tests for two by two tables. For independent binary responses, performs either Fisher's exact test or Blaker's exact test for testing hypotheses about the odds ratio. The commands follow the style of fisher.test, the difference is that for two-sided tests there are three methods for calculating the exact test, and for each of the three methods its matching confidence interval is returned (see details). For paired binary data resulting in a two by two table, performs an exact McNemar’s test.
Usage

```
exact2x2(x, y = NULL, or = 1, alternative = "two.sided",
          tsmethod = NULL, conf.int = TRUE, conf.level = 0.95,
          tol = 0.00001, conditional = TRUE, paired=FALSE,
          plot=FALSE, midp=FALSE)
```

```
fisher.exact(x, y = NULL, or = 1, alternative = "two.sided",
             tsmethod = "minlike", conf.int = TRUE, conf.level = 0.95,
             tol = 0.00001, midp=FALSE)
```

```
blaker.exact(x, y = NULL, or = 1, alternative = "two.sided",
             conf.int = TRUE, conf.level = 0.95, tol = 0.00001)
```

```
mcnemar.exact(x, y=NULL, conf.level=0.9)
```

Arguments

- **x**: either a two-dimensional contingency table in matrix form, or a factor object.
- **y**: a factor object; ignored if x is a matrix.
- **or**: the hypothesized odds ratio. Must be a single numeric.
- **alternative**: indicates the alternative hypothesis and must be one of "two.sided", "greater" or "less". If "two.sided" uses method defined by tsmethod.
- **tsmethod**: one of "minlike", "central", or "blaker". NULL defaults to "minlike" when paired=FALSE and "central" when paired=TRUE or midp=TRUE. Defines type of two-sided method (see details). Ignored if alternative="less" or "greater".
- **conf.int**: logical indicating if a confidence interval should be computed.
- **conf.level**: confidence level for the returned confidence interval. Only used if conf.int = TRUE.
- **tol**: tolerance for confidence interval estimation.
- **conditional**: TRUE. Unconditional exact tests should use `uncondExact2x2`.
- **paired**: logical. TRUE gives exact McNemar’s test, FALSE are all other tests
- **midp**: logical. TRUE gives mid p-values and mid-p CIs. Not supported for tsmethod='minlike’ or 'blaker’
- **plot**: logical. TRUE gives basic plot of point null odds ratios by p-values, for greater plot control use `exact2x2Plot`. Not supported for midp=TRUE.

Details

The motivation for this package is to match the different two-sided conditional exact tests for 2x2 tables with the appropriate confidence intervals.

There are three ways to calculate the two-sided conditional exact tests, motivated by three different ways to define the p-value. The usual two-sided Fisher’s exact test defines the p-value as the sum of probability of tables with smaller likelihood than the observed table (tsmethod="minlike"). The central Fisher’s exact test defines the p-value as twice the one-sided p-values (but with a maximum p-value of 1). Blaker’s (2000) exact test defines the p-value as the sum of the tail probability in the observed tail plus the largest tail probability in the opposite tail that is not greater than the observed tail probability.

In `fisher.test` the p-value uses the two-sample method associated with tsmethod="minlike”, but the confidence interval method associated with tsmethod="central”. The probability that the lower
central confidence limit is less than the true odds ratio is bounded by $1-(1-\text{conf. level})/2$ for
the central intervals, but not for the other two two-sided methods. The confidence intervals in for
exact2x2 match the test associated with alternative. In other words, the confidence interval is the
smallest interval that contains the confidence set that is the inversion of the associated test (see Fay,
2010). The functions fisher.exact and blaker.exact are just wrappers for certain options in
exact2x2.

If $x$ is a matrix, it is taken as a two-dimensional contingency table, and hence its entries should be
nonnegative integers. Otherwise, both $x$ and $y$ must be vectors of the same length. Incomplete cases
are removed, the vectors are coerced into factor objects, and the contingency table is computed from
these.

P-values are obtained directly using the (central or non-central) hypergeometric distribution.

The null of conditional independence is equivalent to the hypothesis that the odds ratio equals one.
‘Exact’ inference can be based on observing that in general, given all marginal totals fixed, the first
element of the contingency table has a non-central hypergeometric distribution with non-centrality
parameter given by the odds ratio (Fisher, 1935). The alternative for a one-sided test is based on the
odds ratio, so alternative = "greater" is a test of the odds ratio being bigger than or.

When paired=TRUE, this denotes there is some pairing of the data. For example, instead of Group
A and Group B, we may have pretest and posttest binary responses. The proper two-sided test for
such a setup is McNemar’s Test, which only uses the off-diagonal elements of the 2x2 table, and
tests that both are equal or not. The exact version is based on the binomial distribution on one of
the off-diagonal values conditioned on the total of both off-diagonal values. We use binom.exact
from the exactci package, and convert the p estimates and confidence intervals (see note) to odds
ratios (see Breslow and Day, 1980, p. 165). The function mcnemar.exact is just a wrapper to call
exact2x2 with paired=TRUE, alternative="two.sided",tmethod="central". One-sided
exact McNemar-type tests may be calculated using the exact2x2 function with paired=TRUE. For
details of McNemar-type tests see Fay (2010, R Journal).

The mid p-value is an adjusted p-value to account for discreteness. The mid-p adjustment is not
guaranteed to give type I error rates that are less than or equal to nominal levels, but gives p-values
that lead to the probability of rejection that is sometimes less than the nominal level and some-
times greater than the nominal level. This adjustment is sometimes used because exact p-values for
discrete data cannot give actual type I error rates equal to the nominal value unless randomization
is done (and that is not typically done because two researchers doing the same method could get
different answers). Essentially, exact p-values lead to the probability of rejecting being less than the
nominal level for most parameter values in the null hypothesis in order to make sure that it is not
greater than the nominal level for ANY parameter values in the null hypothesis. The mid p-value
was studied by Lancaster (1961), and for the 2x2 case by Hirji et al (1991).

Value

A list with class "htest" containing the following components:

- **p.value** the p-value of the test
- **conf.int** a confidence interval for the odds ratio
- **estimate** an estimate of the odds ratio. Note that the conditional Maximum Likelihood
  Estimate (MLE) rather than the unconditional MLE (the sample odds ratio) is used.
- **null.value** the odds ratio under the null, or.
alternative a character string describing the alternative hypothesis
method a character string, changes depending on alternative and tsmethod
data.name a character string giving the names of the data

Note
The default exact confidence intervals for the odds ratio when paired=TRUE (those matching the exact McNemar’s test) are transformations of the Clopper-Pearson exact confidence intervals for a single binomial parameter which are central intervals. See note for binom.exact for discussion of exact binomial confidence intervals.

Author(s)
Michael Fay

References

See Also
fisher.test or mcnemar.test

Examples
```r
## In example 1, notice how fisher.test rejects the null at the 5 percent level,
## but the 95 percent confidence interval on the odds ratio contains 1  
## The intervals do not match the p-value.
## In fisher.exact you get p-values and the matching confidence intervals
example1 <- matrix(c(6,12,12,5),2,2,dimnames=list(c("Group A","Group B"),c("Event","No Event")))
example1
fisher.test(example1)
fisher.exact(example1, tsmethod="minlike")
fisher.exact(example1, tsmethod="central")
blaker.exact(example1)
```
In example 2, this same thing happens, for
tsmethod="minlike"... this cannot be avoided because
of the holes in the confidence set.

```r
example2 <- matrix(c(7,255,30,464),2,2,dimnames=list(c("Group A","Group B"),c("Event","No Event")))
fisher.test(example2)
exact2x2(example2, tsmethod="minlike")
```

You can never get a test-CI inconsistency when tsmethod="central"

```r
exact2x2(example2, tsmethod="central")
```

---

**exact2x2Plot**

Plot p-value function for one 2 by 2 table.

---

**Description**

Plots two-sided p-values as a function of odds ratios. Can plot three types of p-values: the two-sided Fisher’s exact, the central Fisher’s exact (i.e., twice the one-sided Fisher’s exact), and Blaker’s exact.

**Usage**

```r
exact2x2Plot(x, y=NULL, OR = NULL, ndiv = 1000, tsmethod=NULL, 
method = NULL, paired=FALSE, orRange = NULL, dolog = TRUE, 
dolines = FALSE, dopoints = TRUE, doci=TRUE, 
alternative=c("two.sided","less","greater"), 
conf.level=.95, alphaline=TRUE, newplot = TRUE, ...)
```

**Arguments**

- `x` matrix representing the 2 by 2 table
- `y` a factor object; ignored if `x` is a matrix.
- `OR` odds ratio values for plot, if NULL divides orRange into ndiv pieces
- `ndiv` number of pieces to divide up odds ratio range
- `tsmethod` either "minlike", "blaker" or "central"
- `method` same as tsmethod, kept for backward compatibility
- `paired` logical, do paired analysis giving McNemar’s test p-values
- `orRange` range for calculating odds ratios
- `dolog` logical, plot odds ratios on log scale?
- `dolines` logical, add lines to a plot?
- `dopoints` logical, add points to a plot?
- `doci` logical, add vertical lines at confidence interval?
- `alternative` one of "two.sided", "less", "greater", type of alternative for p-values
- `conf.level` when doci=TRUE, level for confidence interval to be plotted
- `alphaline` logical, if doci=TRUE should a line be drawn at the significance level?
- `newplot` logical, start a new plot?
- `...` values passed to plot, points, or lines statement
plotT

Plot or Print ordering function for unconditional exact test

Description

The function orderMat prints the values for the ordering function for all possible values of X1 and X2 in matrix form.

The function plotT plots the ranking of the ordering function on an \( n_1+1 \) by \( n_2+1 \) grid, where each square represents a possible values for \((x_1,x_2)\). The default colors are from dark blue (highest) to light blue to white (middle) to light red to dark red (lowest), with black=NA.

Usage

plotT(x, ...)

## S3 method for class 'function'
plotT(x, n1, n2, delta0 = 1, main = "", ...)

## S3 method for class 'numeric'
plotT(x, n1, n2, delta0 = 1, main = "", ...)

orderMat(x, ...)

## S3 method for class 'function'
orderMat(x, n1, n2, delta0, graphStyle=FALSE, ...)

## S3 method for class 'numeric'
orderMat(x, n1, n2, delta0, graphStyle=FALSE, ...)
power2x2

Arguments

x          object, either a Tstat function, or a vector of all (n1+1)*(n2+1) possible values of the function (see details).
n1          sample size in group 1
n2          sample size in group 2
delta0       null value of parameter (if needed for Tstat function)
main         plot title
graphStyle   logical, order rows with lowest x1 value on the bottom?
...          arguments to be passed to the Tstat function

Details

If x is all the values of the Tstat function, then the values should be ordered by cycling through the x1 values (0 to n1) for each x2 value. Specifically, it should be the result of 
\[ tstatHx1Ln1LxRLnRLdeltaPI \]
where \( X1=\text{rep}(0:n1,n2+1) \) and \( X2=\text{rep}(0:n2,each=n1+1) \).

Examples

```r
parorig<- par(no.readonly=TRUE)
par(mfrow=c(2,2),mar=c(1,3,3,1))
TT1<-pickTstat(method="score", parmtree="ratio", tsmethod="central", alternative="two.sided")
round(orderMat(TT1,8,8,1,graphStyle=TRUE),2)
TT2<-pickTstat(method="simple", parmtree="ratio", tsmethod="central", alternative="two.sided")
TT3<-pickTstat(method="simple", parmtree="difference", tsmethod="central", alternative="two.sided")
plotT(TT2, 8, 8, 1, main="Ratio, Simple")
plotT(TT3, 8, 8, 0, main="Difference, Simple")
plotT(TT1, 8, 8, 1, main="Ratio, Score (delta0=1)")
TF<-pickTstat(method="FisherAdj", parmtree="ratio", tsmethod="central", alternative="two.sided")
plotT(TF, 8, 8, 1, main="FisherAdj")
par(parorig)
```

power2x2          Calculate exact power or sample size for conditional tests for two independent binomials.

Description

Power is calculated by power2x2 which calls exact2x2 function repeatedly. Default (strict=FALSE) does not count rejections in the wrong direction.

Sample size is calculated by ss2x2 which calls power2x2 repeatedly finding the lowest sample size that has at least the nominal power, using the uniroot.integer function from the ssanv package.
Usage

```r
power2x2(p0, p1, n0, n1 = NULL, sig.level = 0.05, alternative = c("two.sided", "one.sided"), paired = FALSE,
          strict = FALSE, tsmethod = NULL, nullOddsRatio = 1,
          errbound = 10^-6, approx = FALSE)
```

```r
ss2x2(p0, p1, power = .80, n1.over.n0 = 1, sig.level = 0.05,
       alternative = c("two.sided", "one.sided"), paired = FALSE,
       strict = FALSE, tsmethod = NULL, nullOddsRatio = 1,
       errbound = 10^-6, print.steps = FALSE, approx = FALSE)
```

Arguments

- `p0`: true event rate in control group
- `p1`: true event rate in treatment group
- `n0`: number of observations in control group
- `n1`: number of observations in treatment group (if NULL n1=n0)
- `sig.level`: significance level (Type I error probability)
- `power`: minimum power for sample size calculation
- `n1.over.n0`: ratio of n1 over n0, allows for non-equal sample size allocation
- `alternative`: character, either "two.sided" or "one.sided", one sided tests the proper direction according to p0 and p1
- `strict`: use strict interpretation of two-sided test, if TRUE counts rejections in wrong direction
- `tsmethod`: two sided method, ignored if strict=FALSE, or alternative equals 'less' or 'greater'. see `exact2x2` for details.
- `nullOddsRatio`: null odds ratio value for tests
- `paired`: logical. TRUE gives power for McNemar's test, FALSE are all other tests (see warning)
- `print.steps`: logical, print steps for calculation of sample size?
- `errbound`: bound on error of calculation
- `approx`: give sample size or power using normal approximation only

Details

Assuming X0 ~ Binomial(n0,p0) and X1 ~ Binomial(n1,p1), calculates the power by repeatedly calling exact2x2 and summing probability of rejection. For speed, the function does not calculate the very unlikely values of X0 and X1 unless errbound=0. Power is exact, but may underestimate by at most errbound.

When strict=FALSE we do not count rejections in the wrong direction. This means that we must know the direction of the rejection, so two.sided tests are calculated as one.sided tests (in the correct direction) with level equal to sig.level/2. This is like using the tsmethod='central'.

When approx=TRUE for power2x2 use a continuity corrected normal approximation (Fleiss, 1981, p. 44). For ss2x2
Value

Both `power2x2` and `ss2x2` return an object of class `power.htest`. A list with elements

- **power**: power to reject
- **n0**: sample size in control group
- **n1**: sample size in treatment group
- **p0**: true event rate in control group
- **p1**: true event rate in treatment group
- **sig.level**: Significance level (Type I error probability)
- **alternative**: alternative hypothesis
- **note**: note about error bound
- **method**: description

Warning

There may be convergence issues using strict=FALSE with tsmethod="minlike" or "blaker" since the power is not guaranteed to be increasing in the sample size.

When paired=TRUE the model for the power calculation is fairly restrictive. It assumes that there is no correlation between the two groups. A better power function is probably needed for this case.

Note

The calculations in ss2x2 can be slow when p0 is close to p1 and/or the power is large. If p0 and p1 are close with large power, it may be safer to first calculate ss2x2 with approx=TRUE to see what the starting value will be close to. If the starting sample sizes are large (>100), it may take a while.

Note when strict=FALSE (default), the two.sided results at the 0.05 level for Fisher’s exact test are like the one.sided Fisher’s exact test at the 0.025 level.

Author(s)

Michael P. Fay

See Also

See `ss.nonadh` function (refinement="Fisher.exact") from the ssanv package for calculation that accounts for nonadherence in proportion of subjects. That function calls `fisher.test`

Examples

```r
power2x2(.2,.8,12,15)
# calculate sample size with 2:1 allocation to groups
ss2x2(.2,.8,n1.over.n0=2,power=.8,approx=TRUE)
ss2x2(.2,.8,n1.over.n0=2,power=.8,print.steps=TRUE)
```
uncondExact2x2

Unconditional exact tests for 2x2 tables

Description

The uncondExact2x2 function tests 2x2 tables assuming two independent binomial responses. Unlike the conditional exact tests which condition on both margins of the 2x2 table (see exact2x2), these unconditional tests only condition on one margin of the 2x2 table (i.e., condition on the sample sizes of the binomial responses). This makes the calculations difficult because now there is a nuisance parameter and calculations must be done over nearly the entire nuisance parameter space.

Usage

uncondExact2x2(x1, n1, x2, n2,
  parmtype = c("difference", "ratio", "oddsratio"), nullparm = NULL,
  alternative = c("two.sided", "less", "greater"),
  conf.int = FALSE, conf.level = 0.95,
  method = c("FisherAdj", "simple", "score", "wald-pooled", "wald-unpooled", "user", "user-fixed"),
  tsmethod = c("central", "square"), midp = FALSE,
  gamma = 0, EplusM=FALSE, tiebreak=FALSE,
  plotprobs = FALSE, control=ucControl(), Tfunc=NULL,...)

uncondExact2x2Pvals(n1, n2, ...)

Arguments

x1 number of events in group 1
n1 sample size in group 1
x2 number of events in group 2
n2 sample size in group 2
parmtype type of parameter of interest, one of "difference", "ratio" or "oddsratio" (see details)
nullparm value of the parameter of interest at null hypothesis, NULL defaults to 0 for parmtype='difference' and 1 for parmtype='ratio' or 'oddsratio'
alternative alternative hypothesis, one of "two.sided", "less", or "greater", default is "two.sided" (see details)
conf.int logical, calculate confidence interval?
conf.level confidence level
method method type, one of "FisherAdj" (default), "simple", "simpleTB", "wald-pooled", "wald-unpooled", "score", "user", or "user-fixed" (see details)
tsmethod two-sided method, either "central" or "square" (see details)
midp logical. Use mid-p-value method?
gamma Beger-Boos adjustment parameter. 0 means no adjustment. (see details).
EplusM logical, do the E+M adjustment? (see details)
tiebreak logical, do tiebreak adjustment? (see details)
plotprobs logical, plot probabilities?
control list of algorithm parameters, see ucControl
Tfunc test statistic function for ordering the sample space when method='user', ignored otherwise (see details)
... extra arguments passed to Tfunc (for uncondExact2x2), or passed to uncondExact2x2Pvals

Details

The uncondExact2x2 function gives unconditional exact tests and confidence intervals for two independent binomial observations. The uncondExact2x2Pvals function repeatedly calls uncondExact2x2 to get the p-values for the entire sample space.

Let X1 be binomial(n1,theta1) and X2 be binomial(n2,theta2). The parmtype determines the parameter of interest: 'difference' is theta2 - theta1, 'ratio' is theta2/theta1, and 'oddsratio' is (theta2*(1-theta1))/(theta1*(1-theta2)).

The options method, parmtype, tsmethod, alternative, EplusM, and tiebreak define some built-in test statistic function, Tstat, that is used to order the sample space, using pickTstat and calcTall. The first 5 arguments of Tstat must be Tstat(x1,n1,x2,n2, delta0), where X1 and X2 must allow vectors, and delta0 is the null parameter value (but delta0 does not need to be used in the ordering). Ordering when parmtype="ratio" or parmtype="oddsratio" is only used when there is information about the parameter. So the ordering function value is not used for ordering when x1=0 and x2=0 for parmtype="ratio", and it is not used when (x1=0 and x2=0) or (x1=n1 and x2=n2) for parmtype="oddsratio".

We describe the ordering functions first for the basic case, the case when tsmethod="central" or alternative!="two.sided", EplusM=FALSE, and tiebreak=FALSE. In this basic case the ordering function, Tstat, is determined by method and parmtype:

- method='simple' - Tstat essentially replaces theta1 with x1/n1 and theta2 with x2/n2 in the parameter definition. If parmtype='difference' then Tstat(x1,n1,x2,n2, delta0) returns x2/n2-x1/n1+delta0. If parmtype='ratio' then the Tstat function returns log(x2/n2) - log(x1/n1) - log(delta0)
  If parmtype='oddsratio' we get log(x2*(N1-x1)/(delta0*x1*(N2-x2))).
- method='wald-pooled' - Tstat is a Z statistic on the difference using the pooled variance (not allowed if parmtype!="difference")
- method='wald-unpooled' - Tstat is a Z statistics on the difference using unpooled variance (not allowed if parmtype!="difference")
- method='score' - Tstat is a Z statistic formed using score statistics, where the parameter is defined by parmtype, and the constrained maximum likelihood estimates of the parameter are calculated by constrMLE.difference, constrMLE.ratio, or constrMLE.oddsratio.
- method='FisherAdj' - Tstat is a one-sided Fisher's 'exact' mid p-value. The mid p-value is an adjustment for ties that technically removes the 'exactness' of the Fisher's p-value...BUT, here we are only using it to order the sample space, so the results of the resulting unconditional test will still be exact.
uncondExact2x2

- method='user' - Tstat is a user supplied statistic given by Tfunc, it must be a function with the first 5 elements of its call being (X1, N1, X2, N2, delta0). The function must returns a vector of length the same as X1 and X2, where higher values suggest larger theta2 compared to theta1 (when tsmethod="square") or higher values suggest more extreme (when tsmethod="square" and alternative="two.sided"). A slower algorithm that does not require monotonicity of one-sided p-values with respect to delta0 is used.

- method='user-fixed' - For advanced users. Tstat is a user supplied statistic given by Tfunc. It should have first 5 elements as described above but its result should not change with delta0 and it must meet Barnard's convexity conditions. If these conditions are met (the conditions are not checked, since checking them will slow the algorithm), then the p-values will be monotonic in delta0 (the null parameter for a two-sided test) and we can use a faster algorithm.

In the basic case, if alternative="two.sided", the argument tsmethod="central" gives the two-sided central method. The p-value is just twice the minimum of the one-sided p-values (or 1 if the doubling is greater than 1).

Now consider cases other than the basic case. The tsmethod="square" option gives the square of the test statistic (when method="simple", "score", "wald-pooled", or "wald-unpooled") and larger values suggest rejection in either direction (unless method='user', then the user supplies any test statistic for which larger values suggest rejection).

The tiebreak=TRUE option breaks ties in a reasonable way when method="simple" (see 'details' section of calcTall). The EplusM=TRUE option performs Lloyd's (2008) E+M ordering on Tstat (see 'details' section of calcTall).

If tiebreak=TRUE and EplusM=TRUE, the tiebreak calculations are always done first.

Berger and Boos (1994) developed a very general method for calculating p-values when a nuisance parameter is present. First, calculate a (1-gamma) confidence interval for the nuisance parameter, check for the supremum over the union of the null hypothesis parameter space and that confidence interval, then add back gamma to the p-value. This adjustment is valid (in other words, applied to exact tests it still gives an adjustment that is exact). The Berger-Boos adjustment is applied when gamma>0.

When method='simple' or method='user-fixed' does a simple grid search algorithm using unirootGrid. No checks are done on the Tstat function when method='user-fixed' to make sure the simple grid search will converge to the proper answer. So method='user-fixed' should be used by advanced users only.

When midp=TRUE the mid p-value is calculated (and the associated confidence interval if conf.int=TRUE) instead of the standard p-value. Loosely speaking, the standard p-value calculates the probability of observing equal or more extreme responses, while the mid p-value calculates the probability of more extreme responses plus 1/2 the probability of equally extreme responses. The tests and confidence intervals when midp=TRUE are not exact, but give type I error rates and coverage of confidence intervals closer to the nominal values. The mid p-value was studied by Lancaster (1961), see vignette on mid p-values for details.

Value

The function uncondExact2x2Pvals returns a (n1+1) by (n2+1) matrix of p-values for all possible x1 and x2 values, while uncondExact2x2 returns a list of class 'htest' with elements:

- statistic proportion in sample 1
uncondExact2x2

- parameter: proportion in sample 2
- p.value: p-value from test
- conf.int: confidence interval on parameter given by parmtyp
- estimate: MLE estimate of parameter given by parmtyp
- null.value: null hypothesis value of parameter given by parmtyp
- alternative: alternative hypothesis
- method: description of test
- data.name: description of data

Warning

The algorithm for calculating the p-values and confidence intervals is based on a series of grid searches. Because the grid searches are often trying to optimize non-monotonic functions, the algorithm is not guaranteed to give the correct answer. At the cost of increasing computation time, better accuracy can be obtained by increasing control$nPgrid, and less often by increasing control$nCIGrid.

Author(s)

Michael P. Fay, Sally A. Hunsberger

References


See Also

See `boschloo` for unconditional exact tests with ordering function based on Fisher's exact p-values.

Examples

```r
# default uses method="FisherAdj"
uncondExact2x2(1,10,9,10,
  parmtyp="ratio")
uncondExact2x2(1,10,9,10,
  method="score",parmtyp="ratio")
```
Calculate power or sample size for any 2x2 test.

Description

The function `Power2x2` and `SS2x2` calculates the power or sample size for any 2x2 test, while the function `uncondPower2x2` calculates power for only tests supported by `uncondExact2x2Pvals`.

Usage

- `Power2x2(n1, n2, theta1, theta2, alpha, pvalFunc, ...)`
- `uncondPower2x2(n1, n2, theta1, theta2, alpha, ...)`
- `SS2x2(theta1, theta2, alpha, pvalFunc, power=0.90, n1.start=10, increaseby=1, n2.over.n1=1, maxiter=50, printSteps=TRUE, ...)`

Arguments

- `n1`: sample size in group 1
- `n2`: sample size in group 2
- `theta1`: probability of success in group 1
- `theta2`: probability of success in group 2
- `alpha`: significance level
- `pvalFunc`: function that inputs x1,n1,x2,n2 and outputs a p-value.
- `power`: target power
- `n1.start`: value of n1 for first iteration
- `increaseby`: positive integer, how much to increase n1 by for each iteration
- `n2.over.n1`: ratio of n2/n1
- `maxiter`: maximum number of iterations
- `printSteps`: logical, should the power and sample size be printed after each iteration?
- `...`: arguments passed to `uncondExact2x2Pvals` (for `uncondPower2x2`), or to `Power2x2` (for `SS2x2`). Not used and saved for future use for `Power2x2`.

Details

The function `Power2x2` is a very simple function to calculate power. It calculates power where rejection is when the p-value from `pvalFunc` is less than or equal to `alpha`. The function `SS2x2` repeatedly calls `Power2x2` as it increases the sample size, stopping when the power is greater than `power`.

The function `uncondPower2x2` is similar except the p-values are calculated by `uncondExact2x2Pvals`. 
unirootGrid

Function to find a root by grid search.

Description

Find the root (value where the function equals 0) of a monotonic function, `func`, using a halving algorithm grid search.

Usage

`unirootGrid(func, power2 = 12, step.up = TRUE, pos.side = FALSE, print.steps = FALSE, power2grid = power2gridRatio, ...)`

Arguments

- `func`: monotonic function
- `power2`: positive integer, number of grid points is $1 + 2^{\text{power2}}$
- `step.up`: logical, start the search at the lower end of the grid and step up?
- `pos.side`: logical, should the root be on the positive side? In other words, should $\text{func}(\text{root}) \geq 0$?


print.steps logical, should each step that is evaluated be printed?
power2grid function that returns the grid. Take one argument, power2
...

Arguments

Details

The grid is defined with the `power2grid` argument that defines a function with an argument `power2`, and returns a grid with $1+2^\text{power2}$ elements. The root is found by a halving algorithm on the grid, so `func` is calculated only $\text{power2}+1$ times. The ‘root’ is the element that is closest to the root, either on the positive side (pos.side=TRUE) or not.

The `unirootGrid` function calls `uniroot.integer` and finds roots based on grid search. The functions `power2gridRatio` and `power2gridDifference` create grids for searching (0,Inf) and (-1,1) respectively. The `power2gridRatio` grid is equally spaced on the log scale with about half of the grid between 0.5 and 2. The function `power2grid` allows more flexibility in defining grids.

Value

A list with elements:

iter number of iterations
f.root value of func at root
root root, element on the grid that is closest to the root on the negative side (if pos.side=FALSE)
bound interval for the accuracy

Author(s)

Michael P. Fay

See Also

`uniroot` and `uniroot.integer`

Examples

```r
# print.steps prints all iterations,
# with x=rank of grid value (e.g., x=1 is lowest value in grid)
# f(x) really is f(grid[x]) where grid is from the power2grid function
unirootGrid(function(x){ x -.37}, power2=10, power2grid=power2gridRatio,
          print.steps=TRUE, pos.side=TRUE)
```
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